# Fusion without diffusions Ryan Chan 

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Divide-and-conquer paradigm
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Divide-and-conquer paradigm

## Divide-and-conquer paradigm

We wish to carry out inferences based on subsets of data and then combine inferences. But why?

- Big Data
- Inference under privacy settings
- Combining views of multiple experts on a topic


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## -Fusion Problem

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Target of interest:

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f(\boldsymbol{x}) \propto f_{1}(\boldsymbol{x}) \cdots f_{C}(\boldsymbol{x})=\prod_{c=1}^{C} f_{c}(\boldsymbol{x})
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## Rejection sampler for $f(x) \propto \sqrt{f_{1} f_{2}}$

Consider $d$-dimensional target density:

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Interested in expectations of the form:

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\begin{aligned}
\mathbb{E}_{\sqrt{f_{1} f_{2}}}[h(X)] & =\mathbb{E}_{\sqrt{f_{1} f_{2}}}[h(X)] \\
& :=\int h(\boldsymbol{x}) \cdot \sqrt{f_{1}(\boldsymbol{x}) \cdot f_{2}(\boldsymbol{x})} \mathrm{d} \boldsymbol{x}
\end{aligned}
$$

## Rejection sampler for $f(x) \propto \sqrt{f_{1} f_{2}}$

We can construct a rejection sampler, since:

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& =\int h \cdot \sqrt{f_{1} \cdot f_{2}} \cdot\left[\frac{f_{1} \vee f_{2}}{f_{1} \vee f_{2}}\right] \cdot\left[\frac{f_{1}+f_{2}}{f_{1}+f_{2}}\right] \mathrm{d} \boldsymbol{x}
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& =\int h \cdot\left[\frac{\sqrt{f_{1} \cdot f_{2}}}{f_{1} \vee f_{2}}\right] \cdot\left[\frac{f_{1} \vee f_{2}}{f_{1}+f_{2}}\right] \cdot\left[f_{1}+f_{2}\right] \mathrm{d} \boldsymbol{x}
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Therefore,

$$
\begin{aligned}
\mathbb{E}_{f}[h(X)] & :=\int h(\boldsymbol{x}) \cdot \sqrt{f_{1}(\boldsymbol{x}) \cdot f_{2}(\boldsymbol{x})} \mathrm{d} \boldsymbol{x} \\
& =\mathbb{E}_{f_{1}+f_{2}}\left[h(X) \cdot\left[\frac{\sqrt{f_{1}(X) \cdot f_{2}(X)}}{f_{1}(X) \vee f_{2}(X)}\right] \cdot\left[\frac{f_{1}(X) \vee f_{2}(X)}{f_{1}(X)+f_{2}(X)}\right]\right]
\end{aligned}
$$

where $x \vee y=\max \{x, y\}, \rho_{1}(X)=\frac{f_{1}(X) \vee f_{2}(X)}{f_{1}(X)+f_{2}(X)} \in[0,1]$ and $\rho_{2}(X)=\frac{\sqrt{f_{1}(X) \cdot f_{2}(X)}}{f_{1}(X) \vee f_{2}(X)} \in\left[\begin{array}{ll}0 & 1\end{array}\right]$

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& =\mathbb{E}_{f_{1}+f_{2}}\left[h(X) \cdot \rho_{1}(X) \cdot \rho_{2}(X)\right] \tag{1}
\end{align*}
$$

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## Rejection sampler for $f(x) \propto \sqrt{f_{1} f_{2}}$

Result: a rejection sampler for $f \propto \sqrt{f_{1} f_{2}}$ with proposal $f_{1}+f_{2}$ and acceptance probability:

$$
\begin{equation*}
\rho_{1}(X) \cdot \rho_{2}(X)=\frac{\sqrt{f_{1}(X) \cdot f_{2}(X)}}{f_{1}(X)+f_{2}(X)} \tag{2}
\end{equation*}
$$

## The $\sqrt{f_{1} f_{2}}$ Algorithm

The algorithm for simulating from $f \propto \sqrt{f_{1} f_{2}}$ proceeds as follows:

1. Simulate $X \sim f_{1}+f_{2}$
2. Accept $X$ with probability $\rho_{1}(X) \cdot \rho_{2}(X)=\frac{\sqrt{f_{1}(X) \cdot f_{2}(X)}}{f_{1}(X)+f_{2}(X)}$, else return to Step 1

Note: can target $f(x) \propto f_{1} f_{2}$ by simply simulating from $f_{1}^{2}$ and $f_{2}^{2}$

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## A simple toy example

Consider $f \propto \sqrt{f_{1} f_{2}}$, where $f_{1} \sim \mathcal{N}(0,1)$ and $f_{2} \sim \mathcal{N}(4,2)$.


## Case where $f_{1}$ and $f_{2}$ are unnormalised

If $f_{1}$ and $f_{2}$ are unnormalised, then simulation from $f_{1}+f_{2}$ is non-trivial
Suppose $y_{1} \sim f_{1}(y)$ and $y_{2} \sim f_{2}(\boldsymbol{y})$. Let $U_{1}, U_{2}, U_{3} \stackrel{i . i . d}{\sim} U[0,1]$ are independent $U[0,1]$ variables. Define the following events:

$$
\begin{aligned}
& \mathcal{F}_{1}=\left\{U_{1}<f_{2}\left(\mu_{1}\right) / f_{1}\left(\varphi_{1}\right)\right\} \\
& \mathcal{F}_{2}=\left\{U_{2}<f_{1}\left(y_{2}\right) / f_{2}\left(y_{2}\right)\right\}
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Case where $f_{1}$ and $f_{2}$ are unnormalised

Define $\boldsymbol{y}^{*}$ and a $0-1$ indicator $/$ as

$$
\left(\boldsymbol{y}^{*}, I\right)= \begin{cases}\left(\boldsymbol{y}^{*}=\boldsymbol{y}_{1}, I=1\right) & \text { if } \overline{\mathcal{F}}_{1} \cap \mathcal{F}_{2} \cap\left\{U_{3} \leq \frac{1}{2}\right\} \text { is true; }  \tag{3}\\ \left(\boldsymbol{y}^{*}=\boldsymbol{y}_{2}, I=1\right) & \text { if } \mathcal{F}_{1} \cap \overline{\mathcal{F}}_{2} \cap\left\{U_{3} \leq \frac{1}{2}\right\} \text { is true; } \\ \left(\boldsymbol{y}^{*}=\boldsymbol{y}_{1}, I=1\right) & \text { if } \mathcal{F}_{1} \cap \mathcal{F}_{2} \cap\left\{U_{3} \leq \frac{1}{2}\right\} \text { is true; } \\ \left(\boldsymbol{y}^{*}=\boldsymbol{y}_{2}, I=1\right) & \text { if } \mathcal{F}_{1} \cap \mathcal{F}_{2} \cap\left\{U_{3}>\frac{1}{2}\right\} \text { is true; } \\ \left(\boldsymbol{y}^{*}=\cdot, I=0\right) & \text { otherwise }\end{cases}
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Conditional on I =1, $y^{*}$ follows the distribution with density
proportional to $f_{1}(\boldsymbol{y})+f_{2}(\boldsymbol{y})$.

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Lemma
Conditional on $I=1, \boldsymbol{y}^{*}$ follows the distribution with density proportional to $f_{1}(\boldsymbol{y})+f_{2}(\boldsymbol{y})$.

## Example

(fill in example for Bayesian Logistic Regression with 7 coefficients)

## Extending to more than 2 sub-posteriors

We can adopt a hierarchical approach to fusion, e.g. for $f \propto f_{1} f_{2} f_{3} f_{4}:$


## Problem of conflicting sub-posteriors

We can rewrite the algorithm for simulating from $f \propto f_{1} f_{2}$ as follows:

1. Simulate $X \sim f_{1}+f_{2}$
2. With probability $\rho_{1}(X)=\frac{\max \left\{f_{1}(X), f_{2}(X)\right\}}{\left(f_{1}(X)+f_{2}(X)\right)}$, continue, else
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3. With probability $p_{2}(X)=\frac{\sqrt{f_{1}(X) f_{2}(X)}}{\max \left\{f_{1}(X), f_{2}(X)\right\}}$, accept $X$ as a sample from $f$, else return to Step 1

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## - Conflicting sub-posteriors

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## Problem of large data-sizes

Algorithm has $\mathcal{O}(n)$ per iteration cost
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& =\sqrt{\frac{f_{1}(X)}{f_{1}(X)+f_{2}(X)}} \cdot \sqrt{\frac{f_{2}(X)}{f_{1}(X)+f_{2}(X)}}
\end{aligned}
$$

## Conclusion and ongoing work

- We've developed a simple rejection sampling algorithm that allows for perfect simulation from $f \propto \sqrt{f_{1} f_{2}}$ by means of simulating from $f_{1}$ and $f_{2}$
- There is ongoing work on Monte Carlo Fusion (which will be spoken about in the next talk by Gareth), which is more suitable for harder cases


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## References

- Dai, H., Pollock, M., and Roberts, G. (2019). Monte Carlo Fusion. Journal of Applied Probability, 56(1):174-191.
- Scott, S. L., Blocker, A. W., Bonassi, F. V., Chipman, H. A., George, E. I., and Mc-Culloch, R. E. (2016). Bayes and big data: The consensus Monte Carlo algorithm. International Journal of Management Science and Engineering Management, 11(2):78-88.

